DETECTION AND ISOLATION OF LEAKAGE AND VALVE FAULTS IN HYDRAULIC SYSTEMS IN VARYING LOADING CONDITIONS, PART 2: FAULT DETECTION AND ISOLATION SCHEME

Jarmo Nurmi and Jouni Mattila

Tampere University of Technology, Department of Intelligent Hydraulics and Automation, P.O. Box 589, 33101 Tampere, Finland

E-mail: jarmo.nurmi@tut.fi

Abstract

Leakages and valve faults are among the most common faults in hydraulic systems. This paper studies the real-time detection and isolation of certain leakage and valve faults based on the results obtained in part one. In the first part, the mathematical model of a hydraulic test bed was analyzed with Global Sensitivity Analysis to facilitate a systematic and verified approach to model-based condition monitoring. In this paper, an Unscented Kalman Filter-based Fault Detection and Isolation scheme for leakage and valve faults of a generic servo valve-controlled hydraulic cylinder is devised. Compared to existing literature, the leakage and valve faults are decoupled from cylinder static and dynamic loading which makes the results generic and applicable to any servo valve-controlled hydraulic cylinder. Moreover, a more comprehensive set of fault patterns for the detection and isolation of leakages and valve faults with experimental and simulation results are presented. We show that detecting an external leakage of as small as 0.17 L/min is possible in some cases, but the accuracy of the method varies considerably. We also report why the isolation of valve faults from leakages is very difficult.

Keywords: fault detection and isolation, leakages, valve faults, varying load, unscented kalman filter, fault patterns

1 Introduction

The idea of model-based condition monitoring is to create a model output $\hat{y}(k)$, which is subtracted from actual measurement $y(k)$ to create a residual $r(k)$ revealing the health of the system (Isermann, 2006). If the model is ideal, the residual remains at zero when the system is operating correctly. But when a fault is introduced, the residual deviates from zero, which is noticed by the fault detection process. Then the fault isolation process takes over and localizes the cause of the fault. The scheme as a whole is called Fault Detection and Isolation (FDI), see Fig. 1.

![Fig. 1: The model-based FDI scheme.](image)

In practice, measurements are noisy and perfect plant models are not possible. Therefore discrepancy between measured and modelled outputs is to be expected. For this reason state estimators (or Kalman Filters) which can consider modelling errors, measurement noise, and utilize measurements to correct model predictions are common in condition monitoring (An et al., 2008; Sepasi et al., 2010).

Previously, An and Sepehri (2008) proposed a method using a fault-free Extended Kalman Filter (EKF) to detect leakages with actuators under unknown external loading. Using the EKF to estimate the external force, they showed that external leakages out of the system and internal leakages across cylinder chambers as small as 0.25 L/min could be detected and isolated. Their approach was proven to work well with sinusoidal and fairly well with pseudorandom inputs.

More recently, Sepasi and Sassani (2010) applied the Unscented Kalman Filter (UKF) to detect leakages and load changes from a hydraulic system with a constant, known external force. They could detect and isolate leakage faults and load changes. However, results were provided using only sinusoidal inputs.

Chen (2010) devised a scheme to detect and isolate internal leakage and sensor offsets. The possibility of decoupling external force from state equations by considering velocity as an input was also proved.

Tan and Sepehri (2002) used the parameters of a nonlinear Volterra model to detect and isolate internal leakage, external leakage, incorrect supply pressure, and contamination in the fluid. Experimental results on the detection of incorrect supply pressure were shown, but the method was offline, which hampers its use for early fault detection. A similar issue affects the fault detection system by Le et al. (1998) where a neural network approach was shown to be sensitive to
relatively high leakages of over 1 L/min.

As opposed to model-based approaches, the use of the wavelet transform by Goharrizi et al. (2010a, 2010b) has produced good results by allowing the detection of an internal leakage of 0.124 L/min. But when external leakages were considered in Goharrizi et al. (2011), it was reported that external leakages of 0.30 L/min could be isolated from an internal leakage of 0.48 L/min, and furthermore external leakages cannot be localized to either side of the actuator, which has been proven to be possible with model-based approaches (An et al., 2008).

In this paper, we extend the methods of Sepasi and Sassani (2010), and An and Sepehri (2008) by treating a more extensive set of faults than those papers and adopting a similar method as Chen (2010) to obtain independence from varying load. The latter is possible as we have a sufficient quality position measurement from which we differentiate velocity, which eliminates the need to estimate them. Thus, we do not need to know the external force nor the load mass, as the information of the mechanism is included in the position and velocity measurement. Therefore, this scheme is more viable in generic hydraulic systems where the load can vary during operation. This paper utilizes the model and Global Sensitivity Analysis (GSA) that was presented for our test bed in (Nurmi and Mattila, 2011). This combines into a systematic approach to model-based condition monitoring compared to the ad hoc approaches currently present. An adaptive threshold is also proposed and experimental results are given with random control signals that are more plausible than sinusoidal inputs. The paper focuses on common leakage and valve faults (Watton, 2007).

This paper is organized as follows. In Section 2, the applicability of the method and the test bed are briefly discussed. In Section 3, the UKF algorithm is introduced and the reduced-order UKF is applied to the test bed. In Section 4, the capability of the UKF and the adaptive threshold scheme are experimentally tested in detecting and isolating leakages and with simulations in detecting and isolating valve faults.

2 Applicability of the Scheme and Test Bed

The FDI scheme used in this paper is applicable to a generic valve-controlled hydraulic cylinder that drives any of the n-DOF manipulator joints affected by any external force and inertia load (Fig. 2). The scheme is considered to be suitable especially for detecting and isolating external and internal leakages.

2.1 Test Bed

As a case study to experimentally validate the scheme, the test bed, in Fig. 3: and Fig. 4:, is used in leakage fault study. It has a 4/3-directional valve that controls the joint cylinder, and three restrictor valves which emulate external leakages between the cylinder and the directional valve (‘External leakage A’ and ‘External leakage B’) and internal leakage across cylinder chambers (‘Internal leakage’), for a list of system components see (Nurmi and Mattila 2011).
3 Unscented Kalman Filter

In this section, an UKF scheme is devised to facilitate model-based FDI. The basis for the scheme originates from Sepasi and Sassani (2010) and An and Sepehri (2008). However, neither scheme is directly applicable to a generic hydraulic system where the load force and mass are not constant or known. Therefore, a modified version is used with decoupling of external force and load mass similar to (Chen, 2010).

This section is organized as follows. In Section 3.1, a generic discrete nonlinear system and its state estimation are introduced. Then in Section 3.2, the UKF algorithm is presented and implemented for the test bed in Section 3.3. Fault detection and isolation principles are discussed in Sections 3.4-3.5.

3.1 Discrete nonlinear system with noise and state estimation

The system is discrete with a nonlinear process $f$ and measurement model $h$ with noise vectors $w$ and $v$:

$$
\begin{align*}
\mathbf{x}_{k+1} &= f(\mathbf{x}_k, \mathbf{u}_k, t_k) + \mathbf{w}_k \\
\mathbf{y}_{k+1} &= h(\mathbf{x}_{k+1}, t_k) + \mathbf{v}_{k+1}
\end{align*}
$$

(1)

where $\mathbf{x}$ is a $N \times 1$ state vector in which $N$ is the number of states, $\mathbf{u}$ is a $U \times 1$ control vector in which $U$ is the number of controls, $t$ is the time, $\mathbf{w}$ is a $N \times 1$ process noise vector, $\mathbf{y}$ is a $M \times 1$ measurement vector in which $M$ is the number of measurements, $\mathbf{v}$ is a $M \times 1$ measurement noise vector and $k$ is a prev. time instant.

The process noise $\mathbf{w}_k$ and measurement noise $\mathbf{v}_{k+1}$ are assumed to be Gaussian ($\mathcal{N}$), white (uncorrelated) and additive with zero mean and covariances $Q_k$ and $R_{k+1}$ with distributions:

$$
\begin{align*}
\mathbf{w}_k &\sim \mathcal{N}(0, Q_k) \\
\mathbf{v}_{k+1} &\sim \mathcal{N}(0, R_{k+1})
\end{align*}
$$

(2)

Hydraulic measurements can be noisy, pressures especially. Considering the noise in the state estimator ensures that residuals are closer to zero in the fault-free situation, hence improving fault detection.

A Kalman-type state estimator for the nonlinear system in Eq. (1) is (Welch and Bishop, 2001):

$$
\begin{align*}
\hat{\mathbf{x}}_{k+1} &= f(\hat{\mathbf{x}}_k, \mathbf{u}_k, t_k) + \mathbf{K}_{k+1}(\mathbf{y}_{k+1} - \hat{\mathbf{y}}_{k+1}) \\
\hat{\mathbf{y}}_{k+1} &= h(\hat{\mathbf{x}}_{k+1}, t_k)
\end{align*}
$$

(3)

where $\hat{\mathbf{x}}$ is the state estimate vector of size $(N - A) \times 1$ with the positive integer $A$ denoting order-reduction. The innovation gain $\mathbf{K}_{k+1}$ is chosen to minimize the mean squared error $E[(\mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1})^2]$. The optimal gain is derived in Simon (2006, pp. 318-320).

Nonlinear state estimation has no optimal solution since the innovation gain is dependent on covariances which are hard to accurately recover after the states are transformed through nonlinear functions. The non-optimal EKF circumvents the problem of nonlinearity by linearizing nonlinear functions around the previous states so that linear estimation techniques from the Kalman Filter (KF) can be applied. However, in the process it introduces approximation errors depending on the severity of the nonlinearity in functions $f$ and $h$.

An approach for tackling the problems of the EKF is the UKF, published by Julier et al. (1995). In (Julier and Uhlmann, 1997; Wan and van der Merwe, 2000) it is shown that the UKF approximates the true mean and covariance of the states more accurately than the EKF with Unscented Transformation (UT). The UT approximates the state distribution with deterministically chosen sigma points assuming that state variables are normally distributed.

Besides the accuracy advantage of UKF over EKF, UKF is also derivative-free, which is useful since calculating and writing long derivatives is error-prone. The given advantages motivate the choice of UKF.

3.2 Unscented Kalman Filter Algorithm

The recursive UKF algorithm can be described in a step by step manner as follows (Wan & van der Merwe 2000):

1. Initialize the filter, Eq. (4)
2. Estimate the a priori state vector $\hat{\mathbf{x}}_{k+1}$ (prediction)
   a. Generate sigma points around the previous estimate, Eq. (5)
   b. Propagate the sigma points through the nonlinear functions, Eq. (6)
   c. Calculate the state mean, Eq. (7)
3. Calculate the a priori error covariance $\mathbf{P}_{k+1}$, Eq. (8)
4. Estimate the a posteriori state vector $\tilde{\mathbf{x}}_{k+1}$
   a. Unscented transformation for measurements (mean and covariance), Eq. (9)
   b. Calculate the cross-covariance between predicted states and measurements, Eq. (10)
   c. Calculate the Kalman gain, Eq. (11)
   d. Update state estimate, Eq. (12)
5. Calculate the a posteriori error covariance $\mathbf{P}_{k+1}$, Eq. (13)
6. Return to step 2
Step 1 is executed once and steps 2-6 are repeated. Steps 2 and 3 constitute the first UT, and step 4a the second. The steps correspond to the following equations:

\[
\begin{align*}
\hat{x}_0 &= E(x_0) \\
P_0 &= E[(x_0 - \hat{x}_0)(x_0 - \hat{x}_0)^T] \\
k &= 0
\end{align*}
\]

\[
\begin{align*}
\hat{x}_k^{(0)} &= \hat{x}_k \\
\hat{x}_k^{(i)} &= \hat{x}_k + \hat{x}_k^{(i)}, \quad i = 1, 2, 3, \ldots, 2N \\
\hat{x}_k^{(i)} &= \left((\sqrt{L + \lambda})P_k\right)^T, \quad i = 1, 2, 3, \ldots, N \\
\hat{x}_k^{(i)} &= -\left((\sqrt{L + \lambda})P_k\right)^T, \quad i = N + 1, \ldots, 2N
\end{align*}
\]

\[
\begin{align*}
\hat{x}_{k+1} &= \sum_{i=0}^{2N} w_i^{(i)}(\hat{x}_{k+1}^{(i)}) + Q_k \\
\hat{y}_{k+1} &= h(\hat{x}_{k+1}, u_k, t_k) \\
P_{xy} &= \sum_{i=0}^{2N}(\hat{y}_{k+1}^{(i)} - \hat{y}_{k+1})(\hat{y}_{k+1}^{(i)} - \hat{y}_{k+1})^T + R_{k+1} \\
P_{xy} &= \sum_{i=0}^{2N}(\hat{x}_{k+1}^{(i)} - \hat{x}_{k+1})(\hat{x}_{k+1}^{(i)} - \hat{x}_{k+1})^T
\end{align*}
\]

where \( E \) is the expectation operator, \( w \) is a weighting coefficient, \( L \) is the dimension of the state vector and \( \lambda \) is a scaling parameter, satisfying \( \lambda = \alpha^2(L + \kappa) - L \). The parameter \( \alpha \) is a tuning factor which determines the spread of the sigma points. A typical value is \( 10^{-1} \). The constant \( \kappa \) is a secondary tuning parameter. Usually it is chosen as zero. The constant \( \beta \) affects the weight of the first error covariance term. An optimal value is \( \beta = 2 \) for normally distributed states. The matrix square root in Eq. (5) should be calculated with Cholesky decomposition for computational efficiency.

If the measurement equations in function \( h \) are linear, steps 4a and 4b can be simplified. The equations in step 4a reduce to (Welch and Bishop, 2001):

\[
\begin{align*}
\hat{y}_{k+1} &= H\hat{x}_{k+1} \\
P_{yy} &= HP_{k+1}H^T + R_{k+1}
\end{align*}
\]

Then step 4b reduces to:

\[
P_{xy} = P_{k+1}H^T
\]

where \( H \) is a measurement matrix of size \( M \times N \).

### 3.3 Unscented Kalman Filter implementation for the test bed

The online estimation of unknown load variables is possible (An et al., 2008), but not very feasible for FDI purposes because the UKF might compensate a fault by incorrectly estimating the load variables, hence making the fault undetectable. The problem is solved with the inclusion of position and velocity measurements to control vector \( u \) (Chen, 2010). The control vector then becomes:

\[
u = [x, \dot{x}, u_c, p_a] = [u_1, u_2, u_3, u_4]
\]

where \( x \) is the position, \( \dot{x} \) is the velocity, \( u_c \) is the valve control signal and \( p_a \) is the supply pressure. The position and velocity measurements could also be included to the state vector for filtering.

In the test bed, boom angle was measured and converted to piston position from which velocity was differentiatated.

#### 3.3.1 Process model

The task of the UKF is to estimate pressures \( p_A \) and \( p_B \), spool position \( x_s \) and spool velocity \( \dot{x}_s \). The reduced-order state vector is thus:

\[
x = [p_A, p_B, x_s, \dot{x}_s]^T = [x_1, x_2, x_3, x_4]^T
\]

Consequently, the discrete-time state space representation from (Nurmi and Mattila, 2011), Eq. (22), reduces to:

\[
\begin{align*}
x_{i+1} &= Aix_i + Bu_i + \xi_i \\
\end{align*}
\]

where \( A_i \) is a state transition matrix, \( B_i \) is a control input matrix, and \( \xi_i \) is a process noise. The measurement equation is:

\[
\begin{align*}
y_i &= Cix_i + Du_i + \zeta_i \\
\end{align*}
\]

In the GSA in (Nurmi and Mattila, 2011), the
effective bulk moduli $B_{e_{fa}}$ and $B_{effB}$ were shown to be somewhat influential in transients, so effort was used to correctly identify them. They were found to be dependent on piston position. In particular, $B_{e_{fa}}$ was quite small when the piston was completely retracted but gradually grew as the piston extended. The following equations taking the flexible volume of the hoses into consideration gave a good approximation:

$$B_{effA} = \frac{B_0B_h(A_Au_1(k) + V_0A)}{(A_Au_1(k) + V_0A)B_h + V_hB_o}$$

$$B_{effB} = \frac{B_0B_h(A_B(x_{max} - u_1(k))) + V_{OB}}{(A_B(x_{max} - u_1(k))) + V_{OB}B_h + V_hB_o}$$

where $B_0$ is the bulk modulus of oil, $B_h$ is the bulk modulus of the hose and $V_h$ is the volume of the hose.

The flow coefficients, shown to be sensitive parameters and treated as constants in the GSA (Nurmi and Mattila, 2011), were not constants but nonlinear functions of spool position. To improve modelling accuracy each flow coefficient was fitted to a third-order polynomial (Muenchhof and Beck, 2008):

$$K_v(x_o) = a_3x_3^3 + a_2x_2^2 + a_1x_1 + a_0$$

Flow coefficient sample points were obtained offline by applying nonlinear parameter estimation techniques to step responses of the valve. Fitting a third-order polynomial to the sample points gave the best compromise between accuracy and complexity; see Root Mean Square Errors (RMSE) in Table 2: The fitted polynomials are shown in Fig. 5: with the polynomial coefficients given in Table 1:

$$\text{Table 1: Flow coefficient polynomials.}$$

<table>
<thead>
<tr>
<th>$K_{vPA}$</th>
<th>$K_{vPB}$</th>
<th>$K_{vAT}$</th>
<th>$K_{vBT}$</th>
</tr>
</thead>
<tbody>
<tr>
<td>(-5.121<em>10^{-8}x_3^3 + 1.556</em>10^{-7}x_3^2 - 1.377<em>10^{-8}x_3 + 2.256</em>10^{-10}m^3/(s*Pa^{1/2})</td>
<td>(9.003<em>10^{-8}x_3^3 + 2.242</em>10^{-7}x_3^2 + 1.73<em>10^{-8}x_3 + 2.921</em>10^{-10}m^3/(s*Pa^{1/2})</td>
<td>(1.031<em>10^{-8}x_3^3 + 2.118</em>10^{-7}x_3^2 + 1.423<em>10^{-8}x_3 + 1.986</em>10^{-10}m^3/(s*Pa^{1/2})</td>
<td>(-2.371<em>10^{-8}x_3^3 + 7.573</em>10^{-7}x_3^2 - 6.588<em>10^{-8}x_3 + 1.91</em>10^{-10}m^3/(s*Pa^{1/2})</td>
</tr>
</tbody>
</table>

In reality, the flow coefficients are also dependent on fluid temperature since the viscosity of the fluid changes with temperature. This modelling was omitted.

### 3.3.2 Initialization

The UKF is initialized as follows:

$$x_0 = [0, 0, 0, 0]^T$$

$$P_0 = \text{diag}([10^{12}, 10^{12}, 10^{-3}, 10^{-3}])$$

$$R = \text{diag}([10^{10}, 10^{10}])$$

$$Q = \text{diag}([10^8, 10^8, 10^{-10}, 10^{-10}])$$

$$\alpha = 0.001$$

$$\beta = 2$$

$$\kappa = 0$$

where $x_o$ is the initial state, $P_0$ is the state covariance matrix, $R$ is the measurement noise matrix and $Q$ is the process noise matrix.

The tuning parameters $\alpha$, $\beta$ and $\kappa$ were chosen according to existing literature (Wan and van der Merwe, 2000). $R$ was chosen to represent measurement noise. The standard deviations of the pressure sensor readings were roughly 0.1 MPa.

The process noise covariance matrix $Q$ represents modeling errors. It proved important to find a balance between process and measurement noise. The variances of process noise were chosen slightly smaller than the variances of measurement noise.

The measurement equations were linear, so the algorithm was reduced according to Eq. (14) and (15). The measurement matrix $H$ was:

$$H = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \end{bmatrix}$$

### 3.4 Fault detection principles

Pressures A and B residuals were calculated for detecting faults, which was justified on the basis of the GSA results provided in (Nurmi and Mattila, 2011). The residuals $r(k)$ were calculated as follows:

$$r(k) = p(k) - \hat{p}(k)$$

where $p(k)$ and $\hat{p}(k)$ are the measured and estimated.
pressures, respectively, and $k$ denotes the current time instant. The residuals were averaged within a moving 5-second window to remove the effect of brief estimation errors. The residual average $\mu_r(k)$ was calculated recursively with (Muenchhof and Isermann, 2005):

$$\mu_r(k) = \mu_r(k-1) + \frac{1}{N}[r(k) - r(k-N)] \quad (24)$$

where $N$, the sample size, was 5000. The recursive formula, although computationally efficient, requires that $N$ samples are stored in memory.

Because the process model cannot be tuned to perfection, a threshold is needed that ensures that the ratio between false alarms (false positives) and undetectable faults (false negatives) is as low as possible. To clarify, the threshold should be constructed in a way that false alarms are minimized, but at the same time the threshold should be as low as possible so that small faults can be detected. Of course to achieve this, the most important factor is the accuracy of the UKF process model.

It is usually enough to use a positive, constant threshold. In this case, the constant positive threshold was not sufficient. A negative threshold was needed so that negative residuals could be used in fault isolation. Also a constant threshold did not work, since the residuals in a faultless situation were larger in chamber A than in B. The reason for this was the load force causing a higher A than B pressure. Consequently, a pressure-dependent threshold was created. A smooth threshold was obtained by averaging both pressures within a moving 5-second window.

Through careful experiments, the following threshold polynomial produced the best results in terms of few false alarms and satisfactory fault detection:

$$t_{\text{pos}}(\mu_p(k)) = 0.012 + 4 \times 10^{-4} \mu_p^2 \quad (25)$$

The unit of the threshold was MPa. The magnitude of the first term was based on the accuracy of the model. The second term ensured that the threshold increased at a suitable rate. The negative threshold was simply $t_{\text{neg}}(\mu_p(k)) = -t_{\text{pos}}(\mu_p(k))$.

### 3.5 Fault isolation patterns

Leakage and valve faults were studied (Fig. 6). The leakage faults were divided into ‘External leakage in chamber A’, ‘External leakage in chamber B’, ‘External leakage in chambers A and B’ and ‘Internal leakage’. The valve faults were divided according to relative opening into ‘Stuck to closed position’, ‘Too small an opening’ and ‘Too large an opening’.

Once a fault was detected, the residuals and other variables were tested against fault patterns that were verified with experiments or simulations to isolate the fault:

**Table 3: Fault patterns for leakage and valve faults.**

<table>
<thead>
<tr>
<th>Fault</th>
<th>$r_{pA}$ ↓</th>
<th>$r_{pB}$ ↓</th>
<th>$r_{pA}$ ↑</th>
<th>$r_{pB}$ ↑</th>
<th>$p_A &gt; p_B$ &amp; $r_{pA}$ ↓</th>
<th>$u_1 &gt; 0$</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>0</td>
<td>1</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>2</td>
<td>0</td>
<td>0</td>
<td>1</td>
<td>0</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>3</td>
<td>1</td>
<td>1</td>
<td>0</td>
<td>1</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>4</td>
<td>1</td>
<td>0</td>
<td>0</td>
<td>1</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>5</td>
<td>1</td>
<td>1</td>
<td>1</td>
<td>0</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>6</td>
<td>1</td>
<td>0</td>
<td>0</td>
<td>1</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>7</td>
<td>1</td>
<td>0</td>
<td>1</td>
<td>0</td>
<td>0</td>
<td>1</td>
</tr>
<tr>
<td>8</td>
<td>1</td>
<td>1</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>9</td>
<td>1</td>
<td>1</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td>1</td>
</tr>
<tr>
<td>10</td>
<td>1</td>
<td>0</td>
<td>0</td>
<td>1</td>
<td>0</td>
<td>0</td>
</tr>
</tbody>
</table>

where ↑ denotes the crossing of the positive threshold, ↓ the crossing of the negative threshold, $r_{pA}$ is the pressure A residual, $r_{pB}$ is the pressure B residual and $u_1$ is the valve control signal.

For example, an external leakage A causes the pressure A residual to cross the negative threshold ($r_{pA} \downarrow = 1$), and the residual $r_{pB}$ to remain within thresholds ($r_{pB} \uparrow = 0$ and $r_{pB} \downarrow = 0$). Considering the direction of the residual we could distinguish simultaneous external leakage A and B from internal leakage and from certain valve faults.

The faults where the valve opens too wide or too little can be *instantly* isolated from internal leakage with a 50-percent probability when both are considered as likely. The possibility of instant detection depends on the test $p_A > p_B$. For example if the first four binaries of the fault code are 0110, and the fifth, the
test ‘\( p_A > p_B \)’ is false, there is offset in the spool position, so either the valve opened too little or too wide. If the test ‘\( u_c > 0 \)’ is true, the valve opening was too small. If it is false, the valve opening was too large. It is possible to isolate the valve fault when the control signal changes from positive to negative or vice versa by observing whether the residuals cross the opposite thresholds. Internal leakage, on the other hand, is not dependent on the sign of valve control signal.

The rationales behind the patterns are as follows. Consider external leakage A as an example. The GSA (Nurmi and Mattila 2011) proved that both pressures are sensitive to a chamber A leakage. However, only the residual \( r_{PA} \) crosses the negative threshold, since the pressure differential \( p_A \) is missing a leakage flow term. The velocity also changes, as shown in the GSA (Nurmi and Mattila 2011). However, its effect to pressure residuals is minor. A similar description applies to external leakage B and internal leakage. However, in internal leakage faults a leakage is present in both chambers. In one chamber the leakage flow is negative, and in the other it is positive.

When the valve is given a positive control signal and it fails to open as much as it should (fault #5), the flow rate to chamber A is too small compared to a fault-free situation. Thus the pressure A measurement is smaller than the UKF estimate and consequently the pressure A residual crosses the negative threshold. At the same time, the pressure B residual crosses the positive threshold because the measured B pressure, as a consequence of the restricting action of the smaller notch BT opening, is larger than the estimated B pressure. Similar explanations apply to faults #6-10. In faults #9-10, the valve is completely closed, so the magnitudes of the residuals reveal the cause.

4 Results

Experimental results for detecting and isolating leakages are given in Section 4.1. In Section 4.2, the detection and isolation of valve faults is studied with simulations.

4.1 Experimental results

The experimental results consist of external leakage A, external leakage B, simultaneous external leakage A and B, and internal leakage. The valve was controlled with fairly random control signals (Fig. 7).
When the pressure A residual approached the threshold, the threshold increased, proving that the thresholds were indeed pressure dependent and so the proposed adaptive threshold worked.

4.1.3 Simultaneous external leakage in chambers A and B

An external leakage A and B were simultaneously introduced to the test bed at the 18th second, as shown in Fig. 11. The external leakage A was detectable only a second later, but the external leakage B took over 30 seconds to detect. The reason for the slow detection was the decreased pressure A that decreased pressure B causing a minor leakage from chamber B. The leakage in chamber B rose to 0.18 L/min before the actual detection of the fault, but a short leakage peak of this magnitude could not be detected. The leakage peaked at 0.52 L/min (average 0.30 L/min) and 0.60 L/min (average 0.34 L/min) at 50 and 65 seconds, and at those instants the threshold of residual B was clearly crossed.

An external leakage A of 0.50 L/min could be detected as that was the leakage magnitude at the time of detection. The leakage averaged at 1.6 L/min between 40 and 70 seconds, but the residual during this period was well over the threshold.

The isolation follows the patterns in Table 3. For the reasons in Section 4.1.1, the residuals were momentarily forced to zero.

4.1.4 Internal leakage

The internal leakage was added to the test bed at the 30th second, as shown in Fig. 12. The positive threshold of pressure B residual was crossed roughly three seconds sooner than the negative of pressure A residual. The leakage varied between 0.35 L/min and 2 L/min, and on average it was 0.94 L/min. During the experiment when the leakage dropped significantly
below the average, the thresholds remained in the fault range, showing that an internal leakage of below 0.50 L/min could be detected, or a leakage in the range of 5-10 % of the flow passing the 4/3-directional valve.

4.2.1 Too small valve opening

The ‘too small valve opening’ fault was introduced to the system at the 35th second, as is illustrated in Fig. 14:

After a second, the fault was detected but could not be immediately isolated since pattern was similar to internal leakage (#3). Once both residuals, as a consequence of the valve control signal change, crossed the other threshold at approximately the 42th second, the fault could be isolated as too small valve opening fault (#5-6).

4.2.2 Too large valve opening

The fault ‘too large valve opening’ (#7-8) was added to the system at the 35th second, as shown in Fig. 15:. The residual behaviour was reversed compared to fault case ‘too small valve opening’. Hence the isolation was possible immediately after detection.

If the internal leakage would have occurred from chamber B to A, this fault could have been isolated, although not instantly following detection.
4.2.3 Valve gets stuck to closed position

At the 35th second, the valve spool got stuck to closed position, shown in Fig. 16:

![Moving average of pressure A residual, opening fault](image1)

![Moving average of pressure B residual, opening fault](image2)

Fig. 16: The valve spool stuck to closed position.

The residual behaviour was similar to internal leakage (#3), except that the residuals crossed the opposite threshold when the valve control signal was reversed at the 40th second. The magnitudes of the residuals revealed the cause as being a closed valve (#9-10).

5 Conclusions and Future Work

In this paper a real-time scheme based on a reduced-order Unscented Kalman Filter (UKF) for detecting and isolating leakage and valve faults from a generic valve-controlled hydraulic cylinder driving a manipulator joint in varying loading conditions was devised and applied to a hydraulic boom test bed. The method is a practical load-independent solution for detecting and isolating especially leakages. In the paper, a comprehensive set of fault patterns were presented, and an adaptive threshold facilitating fault detection was devised. The basis for this work was founded in (Nurmi and Mattila, 2011), where a Global Sensitivity Analysis (GSA) of the test bed was carried out. The results of the analysis were used in this paper, showing that GSA facilitates a systematic and verified approach to model-based condition monitoring. The usefulness of GSA increases with more complicated nonlinear models.

The fault patterns were verified with simulation and experimental studies. The studies, with leakage patterns verified experimentally, showed the possibility of distinguishing external leakage A and B, and simultaneous external leakage A and B from valve faults (spool jamming to too large an opening, to too small an opening and stuck to closed position).

The lowest detectable external leakage was 0.17 L/min, but it varied between experiments. The pressure residuals alone were not enough to distinguish internal leakage and valve faults, so information from the sign of the control signal and the larger chamber pressure were used. The control signal test meant that the time from fault detection to isolation was considerably long in some valve opening or internal leakage fault cases. The fault patterns for those two different faults were found similar, and hence very difficult to distinguish.

The fault-to-isolation time could be shortened and the fault patterns could be expanded by utilizing a spool position measurement. This measurement, however, is not usually available. Moreover, the scheme already requires multiple measurements, and more measurements would increase the probability of sensor failures. The required measurements, however, are: cylinder chamber pressures A and B, supply pressure, valve control signal, piston position or boom angle, and piston velocity or boom angular velocity measurement. If a separate velocity measurement is not available, the velocity could be differentiated from position.

The scheme will be extended to mobile valves using position and velocity sensors more suitable for application domain specific environmental conditions in the future.

Acknowledgement

This work was funded by the Academy of Finland under the project 133273, Sensor network based intelligent condition monitoring of mobile machinery. The authors gratefully acknowledge the Academy of Finland for the financial support.

Nomenclature

<table>
<thead>
<tr>
<th>Symbol</th>
<th>Description</th>
<th>Unit</th>
</tr>
</thead>
<tbody>
<tr>
<td>(\mu_p)</td>
<td>Pressure average</td>
<td>[MPa]</td>
</tr>
<tr>
<td>(r)</td>
<td>Residual</td>
<td>[MPa]</td>
</tr>
<tr>
<td>(t_{pos,neg}(\mu_p))</td>
<td>Positive and negative threshold polynomial</td>
<td>[-]</td>
</tr>
<tr>
<td>(H)</td>
<td>Measurement matrix</td>
<td>[-]</td>
</tr>
<tr>
<td>(K)</td>
<td>Kalman gain</td>
<td>[-]</td>
</tr>
<tr>
<td>(P)</td>
<td>Posteriori state covariance matrix</td>
<td>[-]</td>
</tr>
<tr>
<td>(P^-)</td>
<td>Priori state covariance matrix</td>
<td>[-]</td>
</tr>
<tr>
<td>(P_{xy})</td>
<td>Cross-covariance matrix</td>
<td>[-]</td>
</tr>
<tr>
<td>(P_{yy})</td>
<td>Measurement covariance matrix</td>
<td>[-]</td>
</tr>
<tr>
<td>(Q)</td>
<td>Process noise covariance matrix</td>
<td>[-]</td>
</tr>
<tr>
<td>(R)</td>
<td>Measurement noise covariance matrix</td>
<td>[-]</td>
</tr>
<tr>
<td>(x)</td>
<td>State vector</td>
<td>[-]</td>
</tr>
<tr>
<td>(\hat{x})</td>
<td>State estimate vector</td>
<td>[-]</td>
</tr>
<tr>
<td>(\tilde{x}(i))</td>
<td>Sigma point vector</td>
<td>[-]</td>
</tr>
<tr>
<td>(y)</td>
<td>Measurement vector</td>
<td>[-]</td>
</tr>
</tbody>
</table>
\( v \) Measurement noise vector [-]

\( w \) Process noise vector [-]

Subscripts:
\( k \) Discrete time instant

References


